Nonparametric Methods Lecture 6

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15 Mar. 2010

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Nonparametric Methods Lecture 6 Overview

- Previously, we've assumed that the forms of the underlying densities were of some particular known parametric form.
- But, what if this is not the case?
- Indeed, for most real-world pattern recognition scenarios this assumption is suspect.
- For example, most real-world entities have multimodal distributions whereas all classical parametric densities are unimodal.

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- But, what if this is not the case?
- Indeed, for most real-world pattern recognition scenarios this assumption is suspect.
- For example, most real-world entities have multimodal distributions whereas all classical parametric densities are unimodal.
- We will examine **nonparametric** procedures that can be used with arbitrary distributions and without the assumption that the underlying form of the densities are known.
 - Histograms.
 - Kernel Density Estimation / Parzen Windows.
 - k-Nearest Neighbor Density Estimation.
 - Real Example in Figure-Ground Segmentation

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Histograms



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Histograms



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Histograms



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Histograms



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Histogram Density Representation

• Consider a single continuous variable x and let's say we have a set \mathcal{D} of N of them $\{x_1, \ldots, x_N\}$. Our goal is to model p(x) from \mathcal{D} .

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- Standard histograms simply partition x into distinct bins of width Δ_i and then count the number n_i of observations x falling into bin i.
- To turn this count into a normalized probability density, we simply divide by the total number of observations N and by the width Δ_i of the bins.
- This gives us:

$$p_i = \frac{n_i}{N\Delta_i} \tag{1}$$

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Hence the model for the density p(x) is constant over the width of each bin. (And often the bins are chosen to have the same width Δ_i = Δ.)



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Histogram Density as a Function of Bin Width



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Histogram Density as a Function of Bin Width

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- When Δ is very big (bottom), the resulting density is quite smooth and consequently fails to capture the bimodality of p(x).
- It appears that the *best results* are obtained for some intermediate value of Δ, which is given in the middle figure.
- In principle, a histogram density model is also dependent on the choice of the edge location of each bin.

Analyzing the Histogram Density

• What are the advantages and disadvantages of the histogram density estimator?

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 - Simple to evaluate and simple to use.
 - $\bullet\,$ One can throw away ${\cal D}$ once the histogram is computed.
 - Can be computed sequentially if data continues to come in.

Analyzing the Histogram Density

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- Advantages:
 - Simple to evaluate and simple to use.
 - $\bullet\,$ One can throw away ${\cal D}$ once the histogram is computed.
 - Can be computed sequentially if data continues to come in.
- Disadvantages:
 - The estimated density has discontinuities due to the bin edges rather than any property of the underlying density.
 - Scales poorly (curse of dimensionality): we would have M^D bins if we divided each variable in a D-dimensional space into M bins.

What can we learn from Histogram Density Estimation?

- Lesson 1: To estimate the probability density at a particular location, we should consider the data points that lie within some local neighborhood of that point.
 - This requires we define some distance measure.
 - There is a natural smoothness parameter describing the spatial extent of the regions (this was the bin width for the histograms).

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- Lesson 2: The value of the smoothing parameter should neither be too large or too small in order to obtain good results.

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 - There is a natural smoothness parameter describing the spatial extent of the regions (this was the bin width for the histograms).
- Lesson 2: The value of the smoothing parameter should neither be too large or too small in order to obtain good results.
- With these two lessons in mind, we proceed to kernel density estimation and nearest neighbor density estimation, two closely related methods for density estimation.

- Consider again samples \mathbf{x} from underlying density $p(\mathbf{x})$.
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- Suppose we have n samples $\mathbf{x} \in \mathcal{D}$. The probability of each sample falling into \mathcal{R} is P.
- How will the total number of k points falling into \mathcal{R} be distributed?
- This will be a **binomial distribution**:

$$P_k = \binom{n}{k} P^k (1-P)^{n-k} \tag{3}$$

• The expected value for k is thus

$$\mathcal{E}[k] = nP \tag{4}$$

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- The binomial for k peaks very sharply about the mean. So, we expect k/n to be a very good estimate for the probability P (and thus for the space-averaged density).
- This estimate is increasingly accurate as n increases.



• Assuming continuous $p(\mathbf{x})$ and that \mathcal{R} is so small that $p(\mathbf{x})$ does not appreciably vary within it, we can write:

$$\int_{\mathcal{R}} p(\mathbf{x}') d\mathbf{x}' \simeq p(\mathbf{x}) V \tag{5}$$

where \mathbf{x} is a point within \mathcal{R} and V is the volume enclosed by \mathcal{R} .

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where \mathbf{x} is a point within \mathcal{R} and V is the volume enclosed by \mathcal{R} . • After some rearranging, we get the following estimate for $p(\mathbf{x})$

$$p(\mathbf{x}) \simeq \frac{k}{nV}$$
 (6)

Example

- Simulated an example of example the density at 0.5 for an underlying zero-mean, unit variance Gaussian.
- Varied the volume used to estimate the density.
- Red=1000, Green=2000, Blue=3000, Yellow=4000, Black=5000.



Practical Concerns

- The validity of our estimate depends on two contradictory assumptions:
 - The region R must be sufficiently small the the density is approximately constant over the region.
 - 2 The region \mathcal{R} must be sufficiently large that the number k of points falling inside it is sufficient to yield a sharply peaked binomial.

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- We want p(x), so we need to let V approach 0. However, with a fixed n, R will become so small, that no points will fall into it and our estimate would be useless: p(x) ≃ 0.

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- We want p(x), so we need to let V approach 0. However, with a fixed n, R will become so small, that no points will fall into it and our estimate would be useless: p(x) ≃ 0.
- Note that in practice, we cannot let V to become arbitrarily small because the number of samples is always limited.

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How can we skirt these limitations when an unlimited number of samples if available?

• To estimate the density at x, form a sequence of regions $\mathcal{R}_1, \mathcal{R}_2, \ldots$ containing x with the \mathcal{R}_1 having 1 sample, \mathcal{R}_2 having 2 samples and so on.

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- Let V_n be the volume of \mathcal{R}_n , k_n be the number of samples falling in \mathcal{R}_n , and $p_n(\mathbf{x})$ be the *n*th estimate for $p(\mathbf{x})$:

$$p_n(\mathbf{x}) = \frac{k_n}{nV_n} \tag{7}$$

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• If $p_n(\mathbf{x})$ is to converge to $p(\mathbf{x})$ we need the following three conditions

$$\lim_{n \to \infty} V_n = 0 \tag{8}$$

$$\lim_{n \to \infty} k_n = \infty \tag{9}$$

$$\lim_{n \to \infty} k_n / n = 0 \tag{10}$$

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- There are two common ways of obtaining regions that satisfy these conditions:
 - Shrink an initial region by specifying the volume V_n as some function of n such as $V_n = 1/\sqrt{n}$. Then, we need to show that $p_n(\mathbf{x})$ converges to $p(\mathbf{x})$. (This is like the Parzen window we'll talk about next.)

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 - Specify k_n as some function of n such as k_n = √n. Then, we grow the volume V_n until it encloses k_n neighbors of x. (This is the k-nearest-neighbor).

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Both of these methods converge...

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Parzen Windows

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Parzen Windows

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- The volume of the hypercube is given by

$$V_n = h_n^d \quad . \tag{11}$$

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Parzen Windows

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- The volume of the hypercube is given by

$$V_n = h_n^d \quad . \tag{11}$$

- We can derive an analytic expression for k_n :
 - Define a windowing function:

$$\varphi(\mathbf{u}) = \begin{cases} 1 & |u_j| \le 1/2 & j = 1, \dots, d \\ 0 & \text{otherwise} \end{cases}$$
(12)

- This windowing function φ defines a unit hypercube centered at the origin.
- Hence, φ((x x_i)/h_n) is equal to unity if x_i falls within the hypercube of volume V_n centered at x, and is zero otherwise.

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• The number of samples in this hypercube is therefore given by

$$k_n = \sum_{i=1}^n \varphi\left(\frac{\mathbf{x} - \mathbf{x}_i}{h_n}\right) \quad . \tag{13}$$

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• Substituting in equation (7), $p_n(\mathbf{x}) = k_n/(nV_n)$ yields the estimate

$$p_n(\mathbf{x}) = \frac{1}{n} \sum_{i=1}^n \frac{1}{V_n} \varphi\left(\frac{\mathbf{x} - \mathbf{x}_i}{h_n}\right) \quad . \tag{14}$$

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Hence, the windowing function φ, in this context called a Parzen window, tells us how to weight all of the samples in D to determine p(x) at a particular x.

Example





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Example



• But, what undesirable trait from histograms are inherited by Parzen window density estimates of the form we've just defined?

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Example



- But, what undesirable trait from histograms are inherited by Parzen window density estimates of the form we've just defined?
- Discontinuities...

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Generalizing the Kernel Function

- What if we allow a more general class of windowing functions rather than the hypercube?
- If we think of the windowing function as an interpolator, rather than considering the window function about x only, we can visualize it as a kernel sitting on each data sample x_i in \mathcal{D} .

Generalizing the Kernel Function

- What if we allow a more general class of windowing functions rather than the hypercube?
- If we think of the windowing function as an interpolator, rather than considering the window function about x only, we can visualize it as a kernel sitting on each data sample x_i in D.
- And, if we require the following two conditions on the kernel function φ , then we can be assured that the resulting density $p_n(\mathbf{x})$ will be proper: non-negative and integrate to 1.

$$\varphi(\mathbf{x}) \ge 0$$
 (15)

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$$\int \varphi(\mathbf{u}) d\mathbf{u} = 1 \tag{16}$$

• For our previous case of $V_n = h_n^d$, then it follows $p_n(\mathbf{x})$ will also satisfy these conditions.

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Example: A Univariate Guassian Kernel

• A popular choice of the kernel is the Gaussian kernel:

$$\varphi_h(u) = \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{1}{2}u^2\right] \tag{17}$$



• The resulting density is given by:

$$p(x) = \frac{1}{n} \sum_{i=1}^{n} \frac{1}{h_n \sqrt{2\pi}} \exp\left[-\frac{1}{2h_n^2} (\mathbf{x} - \mathbf{x}_i)^2\right]$$
(18)

 It will give us smoother estimates without the discontinuites from the hypercube kernel.

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Effect of the Window Width Slide I

- An important question is what effect does the window width h_n have on p_n(x)?
- Define $\delta_n(\mathbf{x})$ as

$$\delta_n(\mathbf{x}) = \frac{1}{V_n} \varphi\left(\frac{\mathbf{x}}{h_n}\right) \tag{19}$$

and rewrite $p_n(\mathbf{x})$ as the average

$$p_n(\mathbf{x}) = \frac{1}{n} \sum_{i=1}^n \delta_n(\mathbf{x} - \mathbf{x}_i)$$
(20)

Effect of the Window Width Slide II

• h_n clearly affects both the amplitude and the width of $\delta_n(\mathbf{x})$.



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• But, for any value of h_n , the distribution is normalized:

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• If V_n is too large, the estimate will suffer from too little resolution.

- If V_n is too small, the estimte will suffer from too much variability.
- In theory (with an unlimited number of samples), we can let V_n slowly approach zero as n increases and then $p_n(\mathbf{x})$ will converge to the unknown $p(\mathbf{x})$. But, in practice, we can, at best, seek some compromise.

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Example: Revisiting the Univariate Guassian Kernel



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Example: A Bimodal Distribution



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Parzen Window-Based Classifiers

- Estimate the densities for each category.
- Classify a query point by the label corresponding to the maximum posterior (i.e., one can include priors).

Parzen Window-Based Classifiers

- Estimate the densities for each category.
- Classify a query point by the label corresponding to the maximum posterior (i.e., one can include priors).
- As you guessed it, the decision regions for a Parzen window-based classifier depend upon the kernel function.


Parzen Window-Based Classifiers

- During training, we can made the error arbitrarily low by making the window sufficiently small, but this will have an ill-effect during testing (which is our ultimate need).
- Think of any possibilities for system rules of choosing the kernel?

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- One possibility is to use cross-validation. Break up the data into a training set and a validation set. Then, perform training on the training set with varying bandwidths. Select the bandwidth that minimizes the error on the validation set.

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- One possibility is to use cross-validation. Break up the data into a training set and a validation set. Then, perform training on the training set with varying bandwidths. Select the bandwidth that minimizes the error on the validation set.
- There is little theoretical justification for choosing one window width over another.

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k_n Nearest Neighbor Methods

- Selecting the best window / bandwidth is a severe limiting factor for Parzen window estimators.
- k_n -NN methods circumvent this problem by making the window size a function of the actual training data.

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- Selecting the best window / bandwidth is a severe limiting factor for Parzen window estimators.
- k_n -NN methods circumvent this problem by making the window size a function of the actual training data.
- The basic idea here is to center our window around x and let it grow until it capture k_n samples, where k_n is a function of n.
 - These samples are the k_n nearest neighbors of \mathbf{x} .
 - If the density is high near x then the window will be relatively small leading to good resolution.
 - If the density is low near x, the window will grow large, but it will stop soon after it enters regions of higher density.

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 - If the density is low near x, the window will grow large, but it will stop soon after it enters regions of higher density.
 - In either case, we estimate $p_n(\mathbf{x})$ according to

$$p_n(\mathbf{x}) = \frac{k_n}{nV_n} \tag{22}$$

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$$p_n(\mathbf{x}) = \frac{k_n}{nV_n}$$

• We want k_n to go to infinity as n goes to infinity thereby assuring us that k_n/n will be a good estimate of the probability that a point will fall in the window of volume V_n .

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- Thus, we want k_n/n to go to zero.
- Recall these conditions from the earlier discussion; these will ensure that $p_n(\mathbf{x})$ converges to $p(\mathbf{x})$ as n approaches infinity.

Examples of k_n -NN Estimation

• Notice the discontinuities in the slopes of the estimate.



k-NN Estimation From 1 Sample

- We don't expect the density estimate from 1 sample to be very good, but in the case of *k*-NN it will diverge!
- With n = 1 and $k_n = \sqrt{n} = 1$, the estimate for $p_n(x)$ is

$$p_n(x) = \frac{1}{2|x - x_1|} \tag{23}$$

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But, as we increase the number of samples, the estimate will improve.



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• The k_n -NN Estimator suffers from an analogous flaw from which the Parzen window methods suffer. What is it?

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- We saw earlier that the specification of k_n can lead to radically different density estimates (in practical situations where the number of training samples is limited).

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- But, like the Parzen window size, one choice is as good as another absent any additional information.
- Similarly, in classification scenarios, we can base our judgement on classification error.

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• A reasonable estimate for the posterior is thus

$$P_n(\omega_i | \mathbf{x}) = \frac{p_n(\mathbf{x}, \omega_i)}{\sum_c p_n(\mathbf{x}, \omega_c)} = \frac{k_i}{k}$$
(25)

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 Hence, the posterior probability for ω_i is simply the fraction of samples within the window that are labeled ω_i. This is a simple and intuitive result.

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Example: Figure-Ground Discrimination

Source: Zhao and Davis. Iterative Figure-Ground Discrimination. ICPR 2004.

- Figure-ground discrimination is an important low-level vision task.
- Want to separate the pixels that contain some foreground object (specified in some meaningful way) from the background.



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Example: Figure-Ground Discrimination Source: Zhao and Davis. Iterative Figure-Ground Discrimination. ICPR 2004.

- This paper presents a method for figure-ground discrimination based on non-parametric densities for the foreground and background.
- They use a subset of the pixels from each of the two regions.
- They propose an algorithm called **iterative sampling-expectation** for performing the actual segmentation.
- The required input is simply a region of interest (mostly) containing the object.

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Example: Figure-Ground Discrimination Source: Zhao and Davis. Iterative Figure-Ground Discrimination. ICPR 2004.

- Given a set of n samples $S = {\mathbf{x}_i}$ where each \mathbf{x}_i is a d-dimensional vector.
- We know the kernel density estimate is defined as

$$\hat{p}(\mathbf{y}) = \frac{1}{n\sigma_1 \dots \sigma_d} \sum_{i=1}^n \prod_{j=1}^d \varphi\left(\frac{\mathbf{y}_j - \mathbf{x}_{ij}}{\sigma_j}\right)$$
(26)

where the same kernel φ with different bandwidth σ_j is used in each dimension.

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The Representation Source: Zhao and Davis. Iterative Figure-Ground Discrimination. ICPR 2004.

• The representation used here is a function of RGB:

$$r = R/(R+G+B) \tag{27}$$

$$g = G/(R + G + B) \tag{28}$$

$$s = (R + G + B)/3$$
 (29)

- Separating the chromaticity from the brightness allows them to us a wider bandwidth in the brightness dimension to account for variability due to shading effects.
- And, much narrower kernels can be used on the *r* and *g* chromaticity channels to enable better discrimination.

The Color Density

Source: Zhao and Davis. Iterative Figure-Ground Discrimination. ICPR 2004.

• Given a sample of pixels $S = {\mathbf{x}_i = (r_i, g_i, s_i)}$, the color density estimate is given by

$$\hat{P}(\mathbf{x} = (r, g, s)) = \frac{1}{n} \sum_{i=1}^{n} K_{\sigma_r}(r - r_i) K_{\sigma_g}(g - g_i) K_{\sigma_s}(s - s_i) \quad (30)$$

where we have simplified the kernel definition:

$$K_{\sigma}(t) = \frac{1}{\sigma}\varphi\left(\frac{t}{\sigma}\right) \tag{31}$$

• They use Gaussian kernels

$$K_{\sigma}(t) = \frac{1}{\sqrt{2\pi\sigma}} \exp\left[-\frac{1}{2}\left(\frac{t}{\sigma}\right)^2\right]$$
(32)

with a different bandwidth in each dimension.

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Data-Driven Bandwidth Source: Zhao and Davis. Iterative Figure-Ground Discrimination. ICPR 2004.

• The bandwidth for each channel is calculated directly from the image based on sample statistics.

$$\sigma \approx 1.06\hat{\sigma}n^{-1/5} \tag{33}$$

where $\hat{\sigma}^2$ is the sample variance.

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Initialization: Choosing the Initial Scale Source: Zhao and Davis. Iterative Figure-Ground Discrimination. ICPR 2004.

- For initialization, they compute a distance between the foreground and background distribution by varying the scale of a single Gaussian kernel (on the foreground).
- To evaluate the "significance" of a particular scale, they compute the normalized KL-divergence:

$$\mathsf{nKL}(\hat{P}_{fg}||\hat{P}_{bg}) = \frac{-\sum_{i=1}^{n} \hat{P}_{fg}(\mathbf{x}_i) \log \frac{\hat{P}_{fg}(\mathbf{x}_i)}{\hat{P}_{bg}(\mathbf{x}_i)}}{\sum_{i=1}^{n} \hat{P}_{fg}(\mathbf{x}_i)}$$
(34)

where \hat{P}_{fg} and \hat{P}_{bg} are the density estimates for the foreground and background regions respectively. To compute each, they use about 6% of the pixels (using all of the pixels would lead to quite slow performance).

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Iterative Sampling-Expectation Algorithm Source: Zhao and Davis. Iterative Figure-Ground Discrimination. ICPR 2004.

- Given the initial segmentation, they need to refine the models and labels to adapt better to the image.
- However, this is a chicken-and-egg problem. If we know the labels, we could compute the models, and if we knew the models, we could compute the best labels.

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Iterative Sampling-Expectation Algorithm Source: Zhao and Davis. Iterative Figure-Ground Discrimination. ICPR 2004.

- Given the initial segmentation, they need to refine the models and labels to adapt better to the image.
- However, this is a chicken-and-egg problem. If we know the labels, we could compute the models, and if we knew the models, we could compute the best labels.
- They propose an EM algorithm for this. The basic idea is to alternate between estimating the probability that each pixel is of the two classes, and then given this probability to refine the underlying models.
- EM is guaranteed to converge (but only to a local minimum).

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Initialize using the normalized KL-divergence.

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- Output of the set of pixel from the image to use in the kernel density estimation. This is essentially the 'M' step (because we have a non-parametric density).

- Initialize using the normalized KL-divergence.
- Ouriformly sample a set of pixel from the image to use in the kernel density estimation. This is essentially the 'M' step (because we have a non-parametric density).
- Update the pixel assignment based on maximum likelihood (the 'E' step).
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- Repeat until stable.
 - One can use a hard assignment of the pixels and the kernel density estimator we've discussed, or a soft assignment of the pixels and then a weighted kernel density estimate (the weight is between the different classes).
 - The overall probability of a pixel belonging to the foreground class

$$\hat{P}_{fg}(\mathbf{y}) = \frac{1}{Z} \sum_{i=1}^{n} \hat{P}_{fg}(\mathbf{x}_i) \prod_{j=1}^{d} K\left(\frac{y_j - x_{ij}}{\sigma_j}\right)$$
(35)

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Example: Figure-Ground Discrimination

Results: Stability Source: Zhao and Davis. Iterative Figure-Ground Discrimination. ICPR 2004.



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Results Source: Zhao and Davis. Iterative Figure-Ground Discrimination. ICPR 2004.



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