# Clustering / Unsupervised Methods Lecture 9 

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## Introduction

- Until now, we've assumed our training samples are "labeled" by their category membership.
- Methods that use labeled samples are said to be supervised; otherwise, they're said to be unsupervised.
- However:
- Why would one even be interested in learning with unlabeled samples?
- Is it even possible in principle to learn anything of value from unlabeled samples?


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- Train a classifier on a small set of samples, then tune it up to make it run without supervision on a large, unlabeled set.
- Or, in the reverse direction, let a large set of unlabeled data group automatically, then label the groupings found.


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(9) To find features that will then be useful for categorization.
(5) To gain insight into the nature or structure of the data during the early stages of an investigation.


## Data Clustering

Source: A. K. Jain and R. C. Dubes. Alg. for Clustering Data, Prentiice Hall, 1988.

- What is data clustering?
- Grouping of objects into meaningful categories
- Given a representation of $N$ objects, find $k$ clusters based on a measure of similarity.


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- Natural Classification: degree of similarity among forms.
- Data exploration: discover underlying structure, generate hypotheses, detect anomalies.
- Compression: for organizing data.
- Applications: can be used by any scientific field that collects data!


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- Applications: can be used by any scientific field that collects data!
- Google Scholar: 1500 clustering papers in 2007 alone!


## E.g.: Structure Discovering via Clustering

## Source: http://clusty.com



## E.g.: Topic Discovery

Source: Map of Science, Nature, 2006

- 800,000 scientific papers clustered into 776 topics based on how often the papers were cited together by authors of other papers



## Data Clustering - Formal Definition

- Given a set of $N$ unlabeled examples $D=x_{1}, x_{2}, \ldots, x_{N}$ in a $d$-dimensional feature space, $D$ is partitioned into a number of disjoint subsets $D_{j}$ 's:

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\begin{equation*}
D=\cup_{j=1}^{k} D_{j} \quad \text { where } D_{i} \cup D_{j}=\emptyset, i \neq j \tag{1}
\end{equation*}
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where the points in each subset are similar to each other according to a given criterion $\Phi$.

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- A partition is denoted by

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\begin{equation*}
\pi=\left(D_{1}, D_{2}, \ldots, D_{k}\right) \tag{2}
\end{equation*}
$$

and the problem of data clustering is thus formulated as

$$
\begin{equation*}
\pi^{*}=\underset{\pi}{\operatorname{argmin}} f(\pi), \tag{3}
\end{equation*}
$$

where $f(\cdot)$ is formulated according to $\Phi$.

## $k$-Means Clustering

Source: D. Aurthor and S. Vassilvitskii. $k$-Means++: The Advantages of Careful Seeding

- Randomly initialize $\mu_{1}, \mu_{2}, \ldots, \mu_{c}$
- Repeat until no change in $\mu_{i}$ :
- Classify $N$ samples according to nearest $\mu_{i}$
- Recompute $\mu_{i}$



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First choose $k$ arbitrary centers

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Assign points to closest centers

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## $k$-Means++ Clustering <br> Source: D. Aurthor and S. Vassilvitskii. $k$-Means++: The Advantages of Careful Seeding

- Choose starting centers iteratively.
- Let $D(x)$ be the distance from $x$ to the nearest existing center, take $x$ as new center with probability $\propto D(x)^{2}$.
- Repeat until no change in $\mu_{i}$ :
- Classify $N$ samples according to nearest $\mu_{i}$
- Recompute $\mu_{i}$
- (refer to the slides by D. Author and S. Vassolvitskii for details)


## User's Dilemma

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(5) Which clustering method?
(0) Are the discovered clusters and partition valid?
(1) Does the data have any clustering tendency?

## Cluster Similarity?

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

- Compact Clusters
- Within-cluster distance < between-cluster connectivity
- Connected Clusters
- Within-cluster connectivity > between-cluster connectivity
- Ideal cluster: compact and isolated.



## Representation (features)?

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

- There's no universal representation; they're domain dependent.


nxd pattern matrix

$n \times n$ similarity matrix


## Good Representation

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- A good representation leads to compact and isolated clusters.



## How do we weigh the features?

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

- Two different meaningful groupings produced by different weighting schemes.

http://www.ofai.at/~elias.pampalk/kdd03/animals/


## How do we decide the Number of Clusters?

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

- The samples are generated by 6 independent classes, yet:

ground truth

$k=5$

$k=2$

$k=6$


## Cluster Validity

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

- Clustering algorithms find clusters, even if there are no natural clusters in the data.


100 2D uniform data points

## Comparing Clustering Methods

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

- Which clustering algorithm is the best?



## There's no best Clustering Algorithm!

Source: R. Dubes and A. K. Jain, Clustering Techniques: User's Dilemma, PR 1976

- Each algorithm imposes a structure on data.
- Good fit between model and data $\Rightarrow$ success.


GMM; k=3


Spectral; k=3


GMM; k=2


Spectral; k=2

## Gaussian Mixture Models

- Recall the Gaussian distribution:

$$
\begin{equation*}
\mathcal{N}(x \mid \boldsymbol{\mu}, \boldsymbol{\Sigma})=\frac{1}{(2 \pi)^{d / 2}|\boldsymbol{\Sigma}|^{1 / 2}} \exp \left[-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1}(\mathbf{x}-\boldsymbol{\mu})\right] \tag{4}
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- The Gaussian mixture is a linear superposition of Gaussians in the form:

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\begin{equation*}
p(\mathbf{x})=\sum_{k=1}^{K} \pi_{k} \mathcal{N}\left(\mathbf{x} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}\right) \tag{5}
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- The $\pi_{k}$ are non-negative scalars called mixing coefficients and they govern the relative importance between the various Gaussians in the mixture density. $\sum_{k} \pi_{k}=1$.







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z_{k} & \in\{0,1\}  \tag{6}\\
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- The marginal distribution over $\mathbf{z}$ is specified in terms of the mixing coefficients:

$$
\begin{equation*}
p\left(z_{k}=1\right)=\pi_{k} \tag{8}
\end{equation*}
$$

And, recall, $0 \leq \pi_{k} \leq 1$ and $\sum_{k} \pi_{k}=1$.

- Since $\mathbf{z}$ has a 1-of- $K$ representation, we can also write this distribution as

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\begin{equation*}
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- The conditional distribution of $\mathbf{x}$ given $\mathbf{z}$ is a Gaussian:

$$
\begin{equation*}
p\left(\mathbf{x} \mid z_{k}=1\right)=\mathcal{N}\left(\mathbf{x} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}\right) \tag{10}
\end{equation*}
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or

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\begin{equation*}
p(\mathbf{x} \mid \mathbf{z})=\prod_{k=1}^{K} \mathcal{N}\left(\mathbf{x} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}\right)^{z_{k}} \tag{11}
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- We are interested in the marginal distribution of $\mathbf{x}$ :

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p(\mathbf{x}) & =\sum_{\mathbf{z}} p(\mathbf{x}, \mathbf{z})  \tag{12}\\
& =\sum_{\mathbf{z}} p(\mathbf{z}) p(\mathbf{x} \mid \mathbf{z})  \tag{13}\\
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- So, given our latent variable $\mathbf{z}$, the marginal distribution of $\mathbf{x}$ is a Gaussian mixture.
- If we have $N$ observations $\mathbf{x}_{1}, \ldots, \mathbf{x}_{N}$, then because of our chosen representation, it follows that we have a latent variable $\mathbf{z}_{n}$ for each observed data point $\mathbf{x}_{n}$.


## Component Responsibility Term

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\gamma\left(z_{k}\right) \doteq p\left(z_{k}=1 \mid \mathbf{x}\right) & =\frac{p\left(z_{k}=1\right) p\left(\mathbf{x} \mid z_{k}=1\right)}{\sum_{j=1}^{K} p\left(z_{j}=1\right) p\left(\mathbf{x} \mid z_{j}=1\right)}  \tag{16}\\
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- View $\pi_{k}$ as the prior probability of $z_{k}=1$ and the quantity $\gamma\left(z_{k}\right)$ as the corresponding posterior probability once we have observed $\mathbf{x}$.
- $\gamma\left(z_{k}\right)$ can also be viewed as the responsibility that component $k$ takes for explaining the observation $\mathbf{x}$.


## Sampling from the GMM

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- The figure below-left shows samples from a three-mixture and colors the samples based on their $\mathbf{z}$. The figure below-middle shows samples from the marginal $p(\mathbf{x})$ and ignores $\mathbf{z}$. On the right, we show the $\gamma\left(z_{k}\right)$ for each sampled point, colored accordingly.


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- Similarly, the corresponding latent variables define an $N \times K$ matrix $\mathbf{Z}$ with rows $\mathbf{z}_{n}^{\top}$.


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- Consider this data set as an $N \times d$ matrix $\mathbf{X}$ in which the $n^{\text {th }}$ row is given by $\mathbf{x}_{n}^{\top}$.
- Similarly, the corresponding latent variables define an $N \times K$ matrix $\mathbf{Z}$ with rows $\mathbf{z}_{n}^{\top}$.
- The log-likelihood of the corresponding GMM is given by

$$
\begin{equation*}
\ln p(\mathbf{X} \mid \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})=\sum_{n=1}^{N} \ln \left[\sum_{k=1}^{K} \pi_{k} \mathcal{N}\left(\mathbf{x} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}\right)\right] \tag{18}
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## Maximum-Likelihood

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- Ultimately, we want to find the values of the parameters $\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}$ that maximize this function.
- However, maximizing the log-likelihood terms for GMMs is much more complicated than for the case of a single Gaussian. Why?
- However, maximizing the log-likelihood terms for GMMs is much more complicated than for the case of a single Gaussian. Why?
- The difficulty arises from the sum over $k$ inside of the log-term. The log function no longer acts directly on the Gaussian, and no closed-form solution is available.


## Singularities

- There is a significant problem when we apply MLE to estimate GMM parameters.


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- This term contributes

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\mathcal{N}\left(\mathbf{x}_{n} \mid \mathbf{x}_{n}, \sigma_{j}^{2} \mathbf{I}\right)=\frac{1}{(2 \pi)^{(1 / 2)} \sigma_{j}} \tag{19}
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- Consider the limit $\sigma_{j} \rightarrow 0$ to see that this term goes to infinity and hence the log-likelihood will also go to infinity.
- Thus, the maximization of the log-likelihood function is not a well posed problem because such a singularity will occur whenever one of the components collapses to a single, specific data point.



## Expectation-Maximization for GMMs

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- Recall the conditions that must be satisfied at a maximum of the likelihood function.
- For the mean $\boldsymbol{\mu}_{k}$, setting the derivatives of $\ln p(\mathbf{X} \mid \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})$ w.r.t. $\boldsymbol{\mu}_{k}$ to zero yields

$$
\begin{align*}
0 & =-\sum_{n=1}^{N} \frac{\pi_{k} \mathcal{N}\left(\mathbf{x} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}\right)}{\sum_{j=1}^{K} \pi_{j} \mathcal{N}\left(\mathbf{x} \mid \boldsymbol{\mu}_{j}, \boldsymbol{\Sigma}_{j}\right)} \mathbf{\Sigma}_{k}\left(\mathbf{x}_{n}-\boldsymbol{\mu}_{k}\right)  \tag{20}\\
& =-\sum_{n=1}^{N} \gamma\left(z_{n k}\right) \boldsymbol{\Sigma}_{k}\left(\mathbf{x}_{n}-\boldsymbol{\mu}_{k}\right) \tag{21}
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- Note the natural appearance of the responsibility terms on the RHS.
- Multiplying by $\boldsymbol{\Sigma}_{k}^{-1}$, which we assume is non-singular, gives

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\boldsymbol{\mu}_{k}=\frac{1}{N_{k}} \sum_{n=1}^{N} \gamma\left(z_{n k}\right) \mathbf{x}_{n} \tag{22}
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- We find a similar result for the covariance matrix:

$$
\begin{equation*}
\boldsymbol{\Sigma}_{k}=\frac{1}{N_{k}} \sum_{n=1}^{N} \gamma\left(z_{n k}\right)\left(x_{n}-\boldsymbol{\mu}_{k}\right)\left(x_{n}-\boldsymbol{\mu}_{k}\right)^{\top} \tag{24}
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- We also need to maximize $\ln p(\mathbf{X} \mid \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})$ with respect to the mixing coefficients $\pi_{k}$.
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- Eliminate $\lambda$ and rearrange to obtain:

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\begin{equation*}
\pi_{k}=\frac{N_{k}}{N} \tag{28}
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- Wrong!
- The responsibility terms depend on these parameters in an intricate way:

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\gamma\left(z_{k}\right) \doteq p\left(z_{k}=1 \mid \mathbf{x}\right)=\frac{\pi_{k} \mathcal{N}\left(\mathbf{x} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}\right)}{\sum_{j=1}^{K} \pi_{j} \mathcal{N}\left(\mathbf{x} \mid \boldsymbol{\mu}_{j}, \boldsymbol{\Sigma}_{j}\right)}
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- But, these results do suggest an iterative scheme for finding a solution to the maximum likelihood problem.
(1) Chooce some initial values for the parameters, $\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}$.
(2) Use the current parameters estimates to compute the posteriors on the latent terms, i.e., the responsibilities.
(3) Use the responsibilities to update the estimates of the parameters.
(c) Repeat 2 and 3 until convergence.












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- EM generally tends to take more steps than the K-Means clustering algorithm.
- Each step is more computationally intense than with K-Means too.
- So, one commonly computes K-Means first and then initializes EM from the resulting clusters.
- Care must be taken to avoid singularities in the MLE solution.
- There will generally be multiple local maxima of the likelihood function and EM is not guaranteed to find the largest of these.

Given a GMM, the goal is to maximize the likelihood function with respect to the parameters (the means, the covarianes, and the mixing coefficients).
(1) Initialize the means, $\boldsymbol{\mu}_{k}$, the covariances, $\boldsymbol{\Sigma}_{k}$, and mixing coefficients, $\boldsymbol{\pi}_{k}$. Evaluate the initial value of the log-likelihood.

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\gamma\left(z_{k}\right)=\frac{\pi_{k} \mathcal{N}\left(\mathbf{x} \mid \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}\right)}{\sum_{j=1}^{K} \pi_{j} \mathcal{N}\left(\mathbf{x} \mid \boldsymbol{\mu}_{j}, \boldsymbol{\Sigma}_{j}\right)}
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$$

(3) M-Step Update the parameters using the current responsibilities

$$
\begin{align*}
\boldsymbol{\mu}_{k}^{\text {new }} & =\frac{1}{N_{k}} \sum_{n=1}^{N} \gamma\left(z_{n k}\right) \mathbf{x}_{n}  \tag{29}\\
\boldsymbol{\Sigma}_{k}^{\text {new }} & =\frac{1}{N_{k}} \sum_{n=1}^{N} \gamma\left(z_{n k}\right)\left(\mathbf{x}_{n}-\boldsymbol{\mu}_{k}^{\mathrm{new}}\right)\left(\mathbf{x}_{n}-\boldsymbol{\mu}_{k}^{\mathrm{new}}\right)^{\top}  \tag{30}\\
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where

$$
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(4) Evaluate the log-likelihood

$$
\begin{equation*}
\ln p\left(\mathbf{X} \mid \boldsymbol{\mu}^{\text {new }}, \boldsymbol{\Sigma}^{\text {new }}, \boldsymbol{\pi}^{\text {new }}\right)=\sum_{n=1}^{N} \ln \left[\sum_{k=1}^{K} \pi_{k}^{\text {new }} \mathcal{N}\left(\mathbf{x}_{n} \mid \boldsymbol{\mu}_{k}^{\text {new }}, \boldsymbol{\Sigma}_{k}^{\text {new }}\right)\right] \tag{33}
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$$

(5) Check for convergence of either the parameters of the log-likelihood. If the convergence is not satisfied, set the parameters:

$$
\begin{align*}
\boldsymbol{\mu} & =\boldsymbol{\mu}^{\text {new }}  \tag{34}\\
\boldsymbol{\Sigma} & =\boldsymbol{\Sigma}^{\text {new }}  \tag{35}\\
\boldsymbol{\pi} & =\boldsymbol{\pi}^{\text {new }} \tag{36}
\end{align*}
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and goto step 2.

## A More General View of EM

- The goal of EM is to find maximum likelihood solutions for models having latent variables.


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- Note how the summation over the latent variables appears inside of the log.
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- Note how the summation over the latent variables appears inside of the log.
- Even if the joint distribution $p(\mathbf{X}, \mathbf{Z} \mid \boldsymbol{\theta})$ belongs to the exponential family, the marginal $p(\mathbf{X} \mid \boldsymbol{\theta})$ typically does not.
- If, for each sample $\mathbf{x}_{n}$ we were given the value of the latent variable $\mathbf{z}_{n}$, then we would have a complete data set, $\{\mathbf{X}, \mathbf{Z}\}$, with which maximizing this likelihood term would be straightforward.
- However, in practice, we are not given the latent variables values.
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- Then, in the M -step, we revise the parameters to $\boldsymbol{\theta}^{\text {new }}$ by maximizing this function:

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- Note that the log acts directly on the joint distribution $p(\mathbf{X}, \mathbf{Z} \mid \boldsymbol{\theta})$ and so the M-step maximization will likely be tractable.

